

Truncated unscented particle filter for dealing with non-linear and inequality constraints

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Introduction

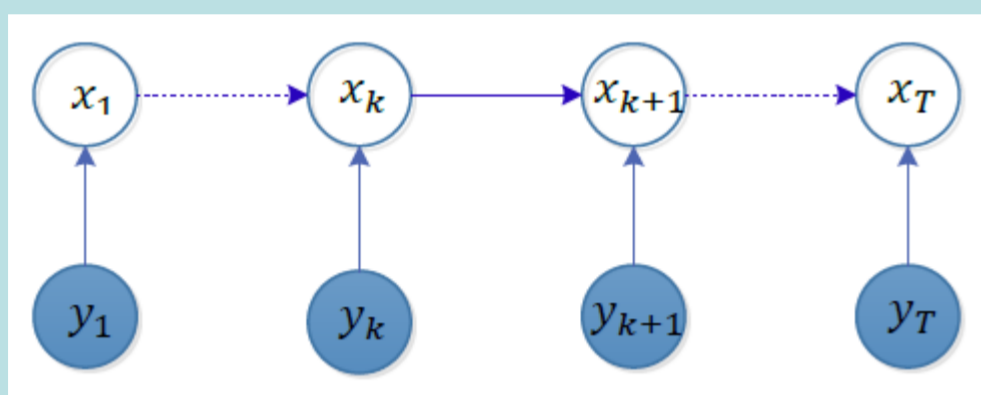
An elegant **truncated unscented particle filtering scheme** considering the provided non-linear and inequality constraint information is proposed:

- A **particle filtering** method is applied to cope with non-linear models and non-Gaussian state distribution.
- A **truncated unscented Kalman filter** is applied as the importance function for sampling new particles.

The advantages of the proposed truncated unscented particle filter algorithm over the state-of-the-art ones are presented by multiple Monte-Carlo simulations.

General constrained tracking problem

Aim: Obtaining the minimum mean square error (MMSE) estimator: $E(x_k|y_{1:k})$



- State model:
 $x_k = f(x_{k-1}, v_k) \sim p(x_k|x_{k-1})$
- Measurement model:
 $y_k = h(x_k, e_k) \sim p(y_k|x_k)$

x_k : state variable

y_k : observation from different types of sensors

For the real state estimation problem, some other information is applied to refine the distribution of the state vector x_k

$$p_C(x_k|z_k) \propto \begin{cases} p(x_k|z_k) & \text{if } x_k \in C_k \\ 0 & \text{otherwise} \end{cases}$$

C_k is the feasible area defined as:

$$C_k = \{x_k | x_k \in R^{n_x}, a_k \leq C_k(x_k) \leq b_k\}$$

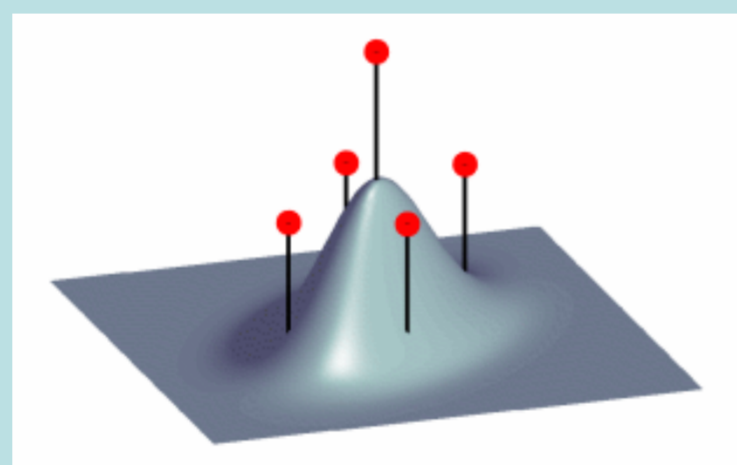
The truncated unscented Kalman filter

The truncated unscented Kalman filter is an extension of the traditional **unscented Kalman filter** by considering the **constraint information**

Initially we have conditional pdf $p_C(x_{k-1}|z_{k-1})$ with mean $\hat{x}_{k-1|k-1}$ and covariance matrix $P_{k-1|k-1}$:

σ -points $\{\chi_{i,k-1|k-1}\}$ and corresponding weights $\{w_{i,k-1|k-1}\}$ are calculated:

$$\begin{aligned} \chi_{0,k-1|k-1} &= \hat{x}_{k-1|k-1} & w_{0,k-1|k-1} &= \frac{\kappa}{n_\chi + \kappa} \\ \chi_{i,k-1|k-1} &= \hat{x}_{k-1|k-1} + (\sqrt{(n_\chi + \kappa)P_{k-1|k-1}})_i \\ w_{i,k-1|k-1} &= \frac{1}{2(n_\chi + \kappa)} \\ \chi_{n_\chi+i,k-1|k-1} &= \hat{x}_{k-1|k-1} - (\sqrt{(n_\chi + \kappa)P_{k-1|k-1}})_i \\ w_{n_\chi+i,k-1|k-1} &= w_{i,k-1|k-1} \end{aligned}$$



According to the σ -points and corresponding weights, the truncated unscented Kalman filter is described as:

Prediction:

$$\begin{aligned} \hat{x}_{k|k-1} &\approx \sum_i w_{i,k-1|k-1} \chi_{i,k|k-1} \\ P_{k|k-1} &\approx \sum_i w_{i,k-1|k-1} (\chi_{i,k|k-1} - \hat{x}_{k|k-1})(\chi_{i,k|k-1} - \hat{x}_{k|k-1})^T + Q_{k-1} \end{aligned}$$

Correction:

$$\begin{aligned} \hat{x}_{k|k} &\approx \hat{x}_{k|k-1} + K_{k|k}(z_k - \hat{z}_{k|k-1}) \\ P_{k|k} &= P_{k|k-1} - K_{k|k}P_{z,k|k-1}K_{k|k}^T \end{aligned}$$

Importance sampling based probability truncation

Truncated distribution after the constraints being considered could be approximated as a Gaussian $N_C(x_k|\hat{x}_{k|k}, P_{k|k})$:

$$\begin{aligned} \hat{x}_{k|k}^c &= \frac{1}{N} \sum_i w_k^{c,i} x_k^{c,i} \\ P_{k|k}^c &= \frac{1}{N} \sum_i w_k^{c,i} (x_k^{c,i} - \hat{x}_{k|k}^c)(x_k^{c,i} - \hat{x}_{k|k}^c)^T \end{aligned}$$

- $x_k^{c,i} \in C_k$ and $x_k^{c,i}$ is sampled from a function $q(x)$
- $w_k^{c,i} \propto N(x_k^{c,i}|\hat{x}_{k|k}, P_{k|k})/q(x_k^{c,i})$

The truncated unscented particle filter

Considering that:

- The posterior distribution may not be accurately represented as a single Gaussian due to constraints
- The nonlinear of the state/measurement models

By taking **the truncated unscented Kalman filter** as the **importance function** for new particles generation, a **truncated unscented particle filtering scheme** is proposed:

Initially, a set of particles and weights $\{x_{k-1}^i, w_{k-1}^i\}_{i=1,\dots,N}$ is applied to approximate $p_C(x_{k-1}|z_{k-1})$

Sampling from the importance function:

Truncated unscented Kalman filter is used to estimate an importance function $N_C(x_k|\hat{x}_{i,k|k}, P_{i,k|k})$ for every particle i

Sampling and rejection:

If the obtained sample x_k^i is within the constraint region, the sample is accepted; otherwise, it is rejected.

Weight calculation:

The weight corresponding to the accepted particle x_k^i is calculated as:

$$w_k^i \propto w_{k-1}^i \frac{p(z_k|x_k^i)p(x_k^i|x_{k-1}^i)}{C_i N_C(x_k^i|\hat{x}_{i,k|k}, P_{i,k|k})}$$

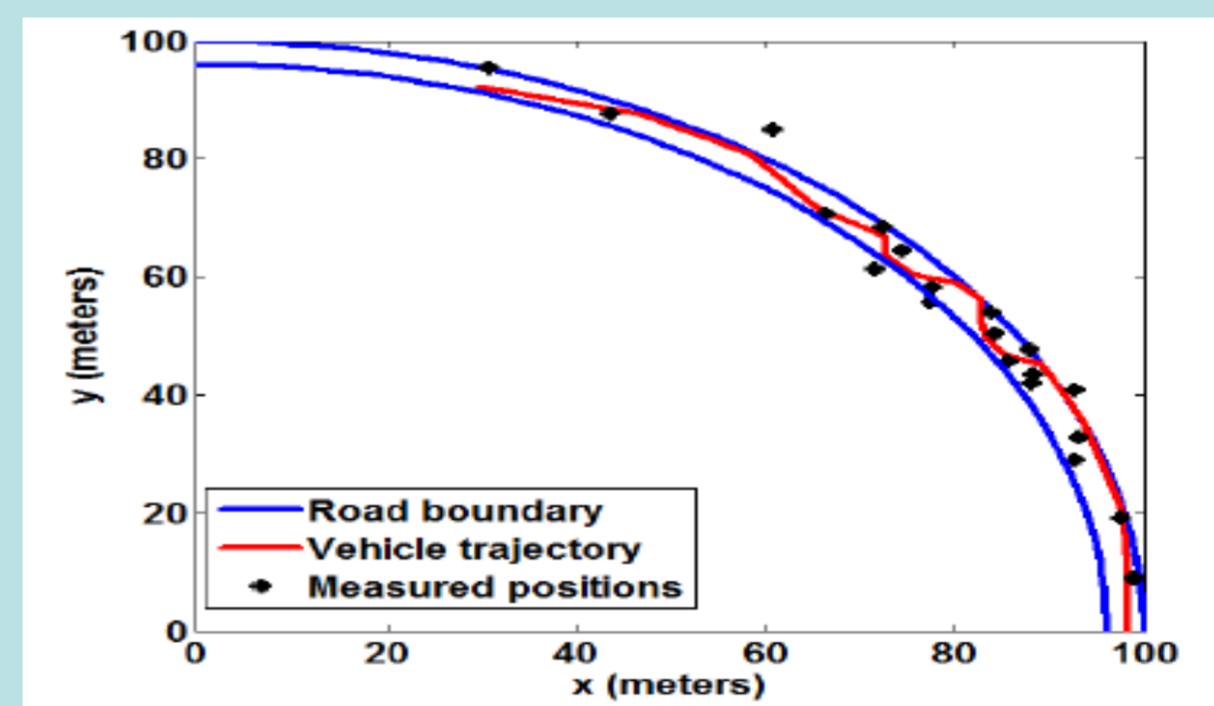
Finally, the weights are summed to one and the state x_t can be estimated as:

$$\hat{x}_k \approx \sum_i w_k^i x_k^i.$$

Simulations

A vehicle is simulated to move with the road constraint:

- The boundaries of the road are defined by two arcs centered at the origin of a Cartesian coordinate system with radius of $r_1 = 96\text{m}$ and $r_2 = 100\text{m}$
- The vehicle dynamics is described by a **constant velocity model**
- **Range** and **bearing angle** are measured



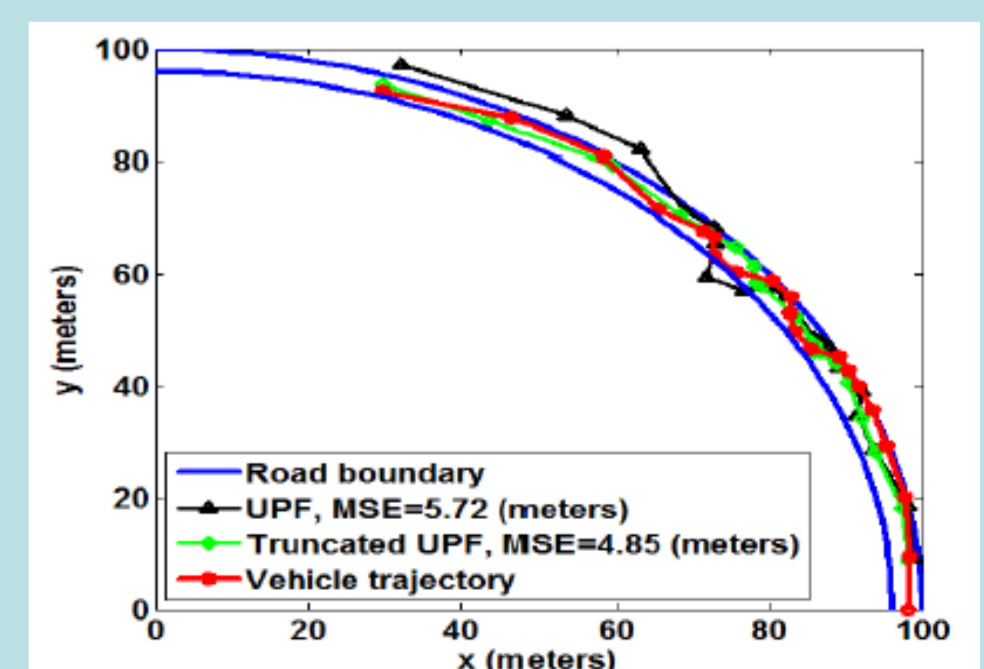
The simulated trajectory of a vehicle moving on a bend road section and the measured positions

Different methods are compared with respect to the mean square errors (MSEs):

The comparison results of the T-UPF and standard UPF

The comparison between different methods (100 times Monte-Carlo simulations)

	Accept-rejection	Projection method	Proposed method
Mean of MSEs (meters)	10.89	6.92	5.40
Standard derivation of MSEs (meters)	7.24	2.33	0.91



Future works

- Considering the 'soft' constraints which concern of probability in different regions
- A more realistic scenario will be considered for the miss detection and false alarms, and the algorithm will be developed under a random finite set framework

